19
Modeling randomized computation

So far we have described randomized algorithms in an informal way, assuming that an operation such as “pick a string \( x \in \{0, 1\}^n \)” can be done efficiently. We have neglected to address two questions:

1. How do we actually efficiently obtain random strings in the physical world?

2. What is the mathematical model for randomized computations, and is it more powerful than deterministic computation?

The first question is of both practical and theoretical importance, but for now let’s just say that there are various physical sources of “random” or “unpredictable” data. A user’s mouse movements and typing pattern, (non solid state) hard drive and network latency, thermal noise, and radioactive decay have all been used as sources for randomness. For example, many Intel chips come with a random number generator built in. One can even build mechanical coin tossing machines (see Fig. 19.1).\(^1\)

In this chapter we focus on the second question: formally modeling probabilistic computation and studying its power. Modeling randomized computation is actually quite easy. We can add the following operations to our NAND, NAND-TM and NAND-RAM programming languages:

\[
\text{foo} = \text{RAND}()
\]

where foo is a variable. The result of applying this operation is that foo is assigned a random bit in \( \{0, 1\} \). (Every time the \text{RAND} operation is invoked it returns a fresh independent random bit.) We call

\(^1\) The output of processes such as above can be thought of as a binary string sampled from some distribution \( \mu \) that might have significant unpredictability (or entropy) but is not necessarily the uniform distribution over \( \{0, 1\}^n \). Indeed, as this paper shows, even (real-world) coin tosses do not have exactly the distribution of a uniformly random string. Therefore, to use the resulting measurements for randomized algorithms, one typically needs to apply a “distillation” or randomness extraction process to the raw measurements to transform them to the uniform distribution.

Figure 19.1: A mechanical coin tosser built for Percy Diaconis by Harvard technicians Steve Sansone and Rick Haggerty
the resulting languages RNAND, RNAND-TM, and RNAND-RAM respectively.

We can use this to define the notion of a function being computed by a randomized \( T(n) \) time algorithm for every nice time bound \( T : \mathbb{N} \rightarrow \mathbb{N} \), as well as the notion of a finite function being computed by a size \( S \) randomized NAND-CIRC program (or, equivalently, a randomized circuit with \( S \) gates that correspond to either the NAND or coin-tossing operations). However, for simplicity we will not define randomized computation in full generality, but simply focus on the class of functions that are computable by randomized algorithms running in polynomial time, which by historical convention is known as BPP:

**Definition 19.1 — BPP.** Let \( F : \{0, 1\}^* \rightarrow \{0, 1\} \). We say that \( F \in \text{BPP} \) if there exist constants \( a, b \in \mathbb{N} \) and an RNAND-RAM program \( P \) such that for every \( x \in \{0, 1\}^* \), on input \( x \), the program \( P \) halts within at most \( a|x|^b \) steps and

\[
\Pr_{r \sim \{0,1\}}[P(x) = F(x)] \geq \frac{2}{3} \tag{19.1}
\]

where this probability is taken over the result of the RAND operations of \( P \). ²

² **BPP** stands for “bounded probability polynomial time”, and is used for historical reasons.

The same polynomial-overhead simulation of NAND-RAM programs by NAND-TM programs we saw in Theorem 12.6 extends to randomized programs as well. Hence the class BPP is the same regardless of whether it is defined via RNAND-TM or RNAND-RAM programs.

### 19.0.1 An alternative view: random coins as an “extra input”

While we presented randomized computation as adding an extra “coin tossing” operation to our programs, we can also model this as being given an additional extra input. That is, we can think of a randomized algorithm \( A \) as a deterministic algorithm \( A' \) that takes two inputs \( x \) and \( r \) where the second input \( r \) is chosen at random from \( \{0, 1\}^m \) for some \( m \in \mathbb{N} \) (see Fig. 19.2). The equivalence to the Definition 19.1 is shown in the following theorem:

**Theorem 19.2 — Alternative characterization of BPP.** Let \( F : \{0, 1\}^* \rightarrow \{0, 1\} \). Then \( F \in \text{BPP} \) if and only if there exists \( a, b \in \mathbb{N} \) and \( G : \{0, 1\}^* \rightarrow \{0, 1\} \) such that \( G \) is in P and for every \( x \in \{0, 1\}^* \),

\[
\Pr_{r \sim \{0,1\}}[G(xr) = F(x)] \geq \frac{2}{3}. \tag{19.2}
\]

**Proof Idea:**

Figure 19.2: The two equivalent views of randomized algorithms. We can think of such an algorithm as having access to an internal RAND() operation that outputs a random independent value in \( \{0, 1\} \) whenever it is invoked, or we can think of it as a deterministic algorithm that in addition to the standard input \( x \in \{0, 1\}^n \) obtains an additional auxiliary input \( r \in \{0, 1\}^m \) that is chosen uniformly at random.
The idea behind the proof is that, as illustrated in Fig. 19.2, we can simply replace sampling a random coin with reading a bit from the extra “random input” \( r \) and vice versa. To prove this rigorously we need to work through some slightly cumbersome formal notation. This might be one of those proofs that is easier to work out on your own than to read.

*  

**Proof of Theorem 19.2.** We start by showing the “only if” direction. Let \( F \in \text{BPP} \) and let \( P \) be an RNAND-RAM program that computes \( F \) as per Definition 19.1, and let \( a, b \in \mathbb{N} \) be such that on every input of length \( n \), the program \( P \) halts within at most \( an^b \) steps. We will construct a polynomial-time NAND-RAM program \( P' \) that computes a function \( G \) satisfying the conditions of (19.2).

The program \( P' \) is very simple:

<table>
<thead>
<tr>
<th>Program ( P' ): (Deterministic NAND-RAM program)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Inputs:</strong></td>
</tr>
<tr>
<td>• ( x \in {0, 1}^n )</td>
</tr>
<tr>
<td>• ( r \in {0, 1}^{an^b} )</td>
</tr>
</tbody>
</table>

| **Goal:** Output \( y \) that has the same distribution as the output of the RNAND-RAM program \( P \) on input \( x \). |

| **Operation:** |
| 1. Copy the string \( r \) to an array \( \text{Coins} \). That is \( \text{Coins}[i] = r_i \) for all \( i \in [an^b] \). |
| 2. Let \( \text{coincounter} \) be a variable and set it to 0. |
| 3. Simulate an execution the RNAND-RAM program \( P \), replacing any line of the form \( \text{foo} = \text{RAND}() \) with the two lines: \( \text{foo} = \text{Coins}[\text{coincounter}] \) and \( \text{coincounter} = \text{coincounter} + 1 \). |

The program \( P' \) is a deterministic polynomial time NAND-RAM program, and so computes some function \( G \in \text{P} \). By its construction, the distribution of \( P'(xr) \) for random \( r \in \{0, 1\}^{2n^b} \) is identical to the distribution of \( P(x) \) (where in the latter case the sample space is the results of the \( \text{RAND}() \) calls), and hence in particular it will hold that \( \Pr_{r \in \{0, 1\}^{an^b}}[P'(xr) = F(x)] \geq 2/3 \).

For the other direction, given a function \( G \in \text{P} \) satisfying the condition (19.2) and a NAND-RAM program \( P' \) that computes \( G \) in polynomial time, we can construct an RNAND-RAM program \( P \) that computes \( F \) in polynomial time. On input \( x \in \{0, 1\}^n \), the program \( P \)
will simply use the RNAND() instruction $a^n b$ times to fill an array $R[0], \ldots, R[a^n b - 1]$ and then execute the original program $P'$ on input $xr$ where $r_i$ is the $i$-th element of the array $R$. Once again, it is clear that if $P'$ runs in polynomial time then so will $P$, and for every input $x$ and $r \in \{0, 1\}^{a^n b}$, the output of $P$ on input $x$ and where the coin tosses outcome is $r$ is equal to $P'(xr)$.

\[\text{Remark 19.3 — Definitions of BPP and NP.} \quad \text{The characterization of BPP Theorem 19.2 is reminiscent of the characterization of NP in Definition 14.1, with the randomness in the case of BPP playing the role of the solution in the case of NP. However, there are important differences between the two:}
\]

- The definition of NP is “one sided”: $F(x) = 1$ if there exists a solution $w$ such that $G(xw) = 1$ and $F(x) = 0$ if for every string $w$ of the appropriate length, $G(xw) = 0$. In contrast, the characterization of BPP is symmetric with respect to the cases $F(x) = 0$ and $F(x) = 1$.

- The relation between NP and BPP is not immediately clear. It is not known whether BPP $\subseteq$ NP, NP $\subseteq$ BPP, or these two classes are incomparable. It is however known (with a non-trivial proof) that if $P = NP$ then BPP $= P$ (see Theorem 19.11).

- Most importantly, the definition of NP is “ineffective,” since it does not yield a way of actually finding whether there exists a solution among the exponentially many possibilities. By contrast, the definition of BPP gives us a way to compute the function in practice by simply choosing the second input at random.

\[\text{Random tapes} \quad \text{Theorem 19.2 motivates sometimes considering the randomness of an RNAND-TM (or RNAND-RAM) program as an extra input. As such, if $A$ is a randomized algorithm that on inputs of length $n$ makes at most $p(n)$ coin tosses, we will often use the notation $A(x;r)$ (where $x \in \{0, 1\}^n$ and $r \in \{0, 1\}^{p(n)}$) to refer to the result of executing $x$ when the coin tosses of $A$ correspond to the coordinates of $r$. This second, or “auxiliary,” input is sometimes referred to as a “random tape.” This terminology originates from the model of randomized Turing machines.}

\[\text{19.0.2 Amplification} \quad \text{The number } 2/3 \text{ might seem arbitrary, but as we’ve seen in Chapter 18 it can be amplified to our liking:}\]
**Theorem 19.4 — Amplification.** Let $P$ be an RNAND-RAM program, $F \in \{0, 1\}^* \rightarrow \{0, 1\}$, and $T : \mathbb{N} \rightarrow \mathbb{N}$ be a nice time bound such that for every $x \in \{0, 1\}^*$, on input $x$ the program $P$ runs in at most $T(|x|)$ steps and moreover $\Pr[P(x) = F(x)] \geq \frac{1}{2} + \epsilon$ for some $\epsilon > 0$. Then for every $k$, there is a program $P'$ taking at most $O(k \cdot T(n)/\epsilon^2)$ steps such that on input $x \in \{0, 1\}^*$, $\Pr[P'(x) = F(x)] > 1 - 2^{-k}$.

**Proof Idea:**
The proof is the same as we’ve seen before in the maximum cut and other examples. We use the Chernoff bound to argue that if we run the program $O(k/\epsilon^2)$ times, each time using fresh and independent random coins, then the probability that the majority of the answers will not be correct will be less than $2^{-k}$. Amplification can be thought of as a “polling” of the choices for randomness for the algorithm (see Fig. 19.3).

Proof of Theorem 19.4. We can run $P$ on input $x$ for $t = 10k/\epsilon^2$ times, using fresh randomness in each execution, and compute the outputs $y_0, \ldots, y_{t-1}$. We output the value $y$ that appeared the largest number of times. Let $X_i$ be the random variable that is equal to 1 if $y_i = F(x)$ and equal to 0 otherwise. The random variables $X_0, \ldots, X_{t-1}$ are i.i.d. and satisfy $\mathbb{E}[X_i] = \Pr[X_i = 1] \geq 1/2 + \epsilon$, and hence by linearity of expectation $\mathbb{E}[\sum_{i=0}^{t-1} X_i] \geq t(1/2 + \epsilon)$. For the plurality value to be incorrect, it must hold that $\sum_{i=0}^{t-1} X_i \leq t/2$, which means that $\sum_{i=0}^{t-1} X_i$ is at least $ct$ far from its expectation. Hence by the Chernoff bound (Theorem 17.11), the probability that the plurality value is not correct is at most $2e^{-\epsilon^2t}$, which is smaller than $2^{-k}$ for our choice of $t$.

There is nothing special about NAND-RAM in Theorem 19.4. The same proof can be used to amplify randomized NAND or NAND-TM programs as well.

**19.1 BPP AND NP COMPLETENESS**

Since “noisy processes” abound in nature, randomized algorithms can be realized physically, and so it is reasonable to propose $\text{BPP}$ rather than $\text{P}$ as our mathematical model for “feasible” or “tractable” computation. One might wonder if this makes all the previous chapters irrelevant, and in particular if the theory of $\text{NP}$ completeness still applies to probabilistic algorithms. Fortunately, the answer is Yes:
Theorem 19.5 — NP hardness and BPP. Suppose that $F$ is NP-hard and $F \in BPP$. Then $NP \subseteq BPP$.

Before seeing the proof, note that Theorem 19.5 implies that if there was a randomized polynomial time algorithm for any NP-complete problem such as $3SAT$, $ISET$ etc., then there would be such an algorithm for every problem in $NP$. Thus, regardless of whether our model of computation is deterministic or randomized algorithms, NP complete problems retain their status as the “hardest problems in $NP$.”

Proof Idea:

The idea is to simply run the reduction as usual, and plug it into the randomized algorithm instead of a deterministic one. It would be an excellent exercise, and a way to reinforce the definitions of NP-hardness and randomized algorithms, for you to work out the proof for yourself. However for the sake of completeness, we include this proof below.

⋆

Proof of Theorem 19.5. Suppose that $F$ is NP-hard and $F \in BPP$. We will now show that this implies that $NP \subseteq BPP$. Let $G \in NP$.

By the definition of NP-hardness, it follows that $G \leq_p F$, or that in other words there exists a polynomial-time computable function $R : \{0, 1\}^* \rightarrow \{0, 1\}^*$ such that $G(x) = F(R(x))$ for every $x \in \{0, 1\}^*$. Now if $F$ is in $BPP$ then there is a polynomial-time RNAND-TM program $P$ such that

$$\Pr[P(y) = F(y)] \geq 2/3$$

(19.3)

for every $y \in \{0, 1\}^*$ (where the probability is taken over the random coin tosses of $P$). Hence we can get a polynomial-time RNAND-TM program $P'$ to compute $G$ by setting $P'(x) = P(R(x))$. By (19.3) $\Pr[P'(x) = F(R(x))] \geq 2/3$ and since $F(R(x)) = G(x)$ this implies that $\Pr[P'(x) = G(x)] \geq 2/3$, which proves that $G \in BPP$.

Most of the results we’ve seen about NP hardness, including the search to decision reduction of Theorem 15.1, the decision to optimization reduction of Theorem 15.2, and the quantifier elimination result of Theorem 15.5, all carry over in the same way if we replace $P$ with BPP as our model of efficient computation. Thus if $NP \subseteq BPP$ then we get essentially all of the strange and wonderful consequences of $P = NP$. Unsurprisingly, we cannot rule out this possibility. In fact, unlike $P = EXP$, which is ruled out by the time hierarchy theorem, we don’t even know how to rule out the possibility that $BPP = EXP$! Thus a priori it’s possible (though seems highly unlikely) that randomness is a magical tool that allows us to speed up arbitrary exponential time
computation. Nevertheless, as we discuss below, it is believed that randomization’s power is much weaker and BPP lies in much more “pedestrian” territory.

19.2 THE POWER OF RANDOMIZATION

A major question is whether randomization can add power to computation. Mathematically, we can phrase this as the following question: does BPP = P? Given what we’ve seen so far about the relations of other complexity classes such as P and NP, or NP and EXP, one might guess that:

1. We do not know the answer to this question.

2. But we suspect that BPP is different than P.

One would be correct about the former, but wrong about the latter. As we will see, we do in fact have reasons to believe that BPP = P. This can be thought of as supporting the extended Church Turing hypothesis that deterministic polynomial-time NAND-TM program (or, equivalently, polynomial-time Turing machines) capture what can be feasibly computed in the physical world.

We now survey some of the relations that are known between BPP and other complexity classes we have encountered. (See also Fig. 19.4.)

19.2.1 Solving BPP in exponential time

It is not hard to see that if F is in BPP then it can be computed in exponential time.

\textbf{Theorem 19.6} — Simulating randomized algorithms in exponential time.

\[ \text{BPP} \subseteq \text{EXP} \]

The proof of Theorem 19.6 readily follows by enumerating over all the (exponentially many) choices for the random coins. We omit the formal proof, as doing it by yourself is an excellent way to get comfortable with Definition 19.1.

19.2.2 Simulating randomized algorithms by circuits or straight-line programs.

We have seen in Theorem 12.12 that if F is in P, then there is a polynomial \( p : \mathbb{N} \to \mathbb{N} \) such that for every \( n \), the restriction \( F_{|n} \) of \( F \) to inputs \( \{0,1\}^n \) is in \( \text{SIZE}(p(n)) \). (In other words, that \( P \subseteq \text{P}^{\text{poly}} \).) A priori it is
not at all clear that the same holds for a function in \textbf{BPP}, but this does turn out to be the case.

\textbf{Theorem 19.7 --- Randomness does not help for non uniform computation.}

\textbf{BPP} \subseteq \textbf{P}/\text{poly}. That is, for every \( F \in \text{BPP} \), there exist some \( a, b \in \mathbb{N} \) such that for every \( n > 0 \), \( F_{n} \in \text{SIZE}(an^{b}) \) where \( F_{n} \) is the restriction of \( F \) to inputs in \( \{0,1\}^{n} \).

\textbf{Proof Idea:}

The idea behind the proof is that we can first amplify by repetition the probability of success from \( 2/3 \) to \( 1 - 0.1 \cdot 2^{-n} \). This will allow us to show that there exists a single fixed choice of “favorable coins” that would cause the algorithm to output the right answer on \textit{all} of the possible \( 2^{n} \) inputs. We can then use the standard “unravelling the loop” technique to transform an RNAND-TM program to an RNAND-CIRC program, and “hardwire” the favorable choice of random coins to transform the RNAND-CIRC program into a plain old deterministic NAND-CIRC program.

\* 

\textbf{Proof of Theorem 19.7.} Suppose that \( F \in \text{BPP} \). Let \( P \) be a polynomial-time RNAND-TM program that computes \( F \) as per Definition 19.1. Using Theorem 19.4, we can amplify the success probability of \( P \) to obtain an RNAND-TM program \( P' \) that is at most a factor of \( O(n) \) slower (and hence still polynomial time) such that for every \( x \in \{0,1\}^{n} \)

\[
\Pr_{r \sim \{0,1\}^{m}}[P'(x; r) = F(x)] \geq 1 - 0.1 \cdot 2^{-n}, \quad (19.4)
\]

where \( m \) is the number of coin tosses that \( P' \) uses on inputs of length \( n \). We use the notation \( P'(x; r) \) to denote the execution of \( P' \) on input \( x \) and when the result of the coin tosses corresponds to the string \( r \).

For every \( x \in \{0,1\}^{n} \), define the “bad” event \( B_{x} \) to hold if \( P'(x) \neq F(x) \), where the sample space for this event consists of the coins of \( P' \). Then by \( (19.4) \), \( \Pr[B_{x}] \leq 0.1 \cdot 2^{-n} \) for every \( x \in \{0,1\}^{n} \). Since there are \( 2^{n} \) many such \( x \)’s, by the union bound we see that the probability that the \textit{union} of the events \( \{B_{x}\}_{x \in \{0,1\}^{n}} \) is at most 0.1. This means that if we choose \( r \sim \{0,1\}^{m} \), then with probability at least 0.9 it will be the case that for every \( x \in \{0,1\}^{n} \), \( F(x) = P'(x; r) \). (Indeed, otherwise the event \( B_{x} \) would hold for some \( x \).) In particular, because of the mere fact that the the probability of \( \cup_{x \in \{0,1\}^{n}} B_{x} \) is smaller than 1, this means that \textit{there exists} a particular \( r^{*} \in \{0,1\}^{m} \) such that

\[
P'(x; r^{*}) = F(x) \quad (19.5)
\]
for every $x \in \{0, 1\}^n$.

Now let us use the standard “unravelling the loop” the technique and transform $P'$ into a NAND-CIRC program $Q$ of polynomial in $n$ size, such that $Q(xr) = P'(x; r)$ for every $x \in \{0, 1\}^n$ and $r \in \{0, 1\}^m$. Then by “hardwiring” the values $r_0, \ldots, r_{m-1}$ in place of the last $m$ inputs of $Q$, we obtain a new NAND-CIRC program $Q_r$ that satisfies by (19.5) that $Q_r(x) = F(x)$ for every $x \in \{0, 1\}^n$. This demonstrates that $F'_{\infty}$ has a polynomial sized NAND-CIRC program, hence completing the proof of Theorem 19.7.

**Remark 19.8 — Randomness and non uniformity.** The proof of Theorem 19.7 actually yields more than its statement. We can use the same “unrolling the loop” arguments we’ve used before to show that the restriction to $\{0, 1\}^n$ of every function in BPP is also computable by a polynomial-size RNAND-CIRC program (i.e., NAND-CIRC program with the RAND operation). Like in the P vs SIZE(poly(n)) case, there are also functions outside BPP whose restrictions can be computed by polynomial-size RNAND-CIRC programs. Nevertheless the proof of Theorem 19.7 shows that even such functions can be computed by polynomial sized NAND-CIRC programs without using the rand operations. This can be phrased as saying that $BPSIZE(T(n)) \subseteq SIZE(O(nT(n)))$ (where $BPSIZE$ is defined in the natural way using RNAND programs). The stronger version of Theorem 19.7 we mentioned can be phrased as saying that $BPP/poly = P/poly$.

### 19.3 DERANDOMIZATION

The proof of Theorem 19.7 can be summarized as follows: we can replace a poly(n)-time algorithm that tosses coins as it runs with an algorithm that uses a single set of coin tosses $r^* \in \{0, 1\}^{poly(n)}$ which will be good enough for all inputs of size $n$. Another way to say it is that for the purposes of computing functions, we do not need “online” access to random coins and can generate a set of coins “offline” ahead of time, before we see the actual input.

But this does not really help us with answering the question of whether $BPP$ equals $P$, since we still need to find a way to generate these “offline” coins in the first place. To derandomize an RNAND-TM program we will need to come up with a single deterministic algorithm that will work for all input lengths. That is, unlike in the case of RNAND-CIRC programs, we cannot choose for every input length $n$ some string $r^* \in \{0, 1\}^{poly(n)}$ to use as our random coins.
Can we derandomize randomized algorithms, or does randomness add an inherent extra power for computation? This is a fundamentally interesting question but is also of practical significance. Ever since people started to use randomized algorithms during the Manhattan project, they have been trying to remove the need for randomness and replace it with numbers that are selected through some deterministic process. Throughout the years this approach has often been used successfully, though there have been a number of failures as well.

A common approach people used over the years was to replace the random coins of the algorithm by a “randomish looking” string that they generated through some arithmetic progress. For example, one can use the digits of \( \pi \) for the random tape. Using these type of methods corresponds to what von Neumann referred to as a “state of sin”. (Though this is a sin that he himself frequently committed, as generating true randomness in sufficient quantity was and still is often too expensive.) The reason that this is considered a “sin” is that such a procedure will not work in general. For example, it is easy to modify any probabilistic algorithm \( A \) such as the ones we have seen in Chapter 18, to an algorithm \( A' \) that is guaranteed to fail if the random tape happens to equal the digits of \( \pi \). This means that the procedure “replace the random tape by the digits of \( \pi \)” does not yield a general way to transform a probabilistic algorithm to a deterministic one that will solve the same problem. Of course, this procedure does not always fail, but we have no good way to determine when it fails and when it succeeds. This reasoning is not specific to \( \pi \) and holds for every deterministically produced string, whether it obtained by \( \pi \), \( e \), the Fibonacci series, or anything else.

An algorithm that checks if its random tape is equal to \( \pi \) and then fails seems to be quite silly, but this is but the “tip of the iceberg” for a very serious issue. Time and again people have learned the hard way that one needs to be very careful about producing random bits using deterministic means. As we will see when we discuss cryptography, many spectacular security failures and break-ins were the result of using “insufficiently random” coins.

### 19.3.1 Pseudorandom generators

So, we can’t use any single string to “derandomize” a probabilistic algorithm. It turns out however, that we can use a collection of strings to do so. Another way to think about it is that rather than trying to eliminate the need for randomness, we start by focusing on reducing the amount of randomness needed. (Though we will see that if we reduce the randomness sufficiently, we can eventually get rid of it altogether.)

We make the following definition:
**Definition 19.9 — Pseudorandom generator.** A function $G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m$ is a $(T, \epsilon)$-pseudorandom generator if for every NAND-CIRC program $P$ with $m$ inputs and one output of at most $T$ lines,

\[
\left| \Pr_{s \sim \{0, 1\}^\ell} [P(G(s)) = 1] - \Pr_{r \sim \{0, 1\}^m} [P(r) = 1] \right| < \epsilon
\] (19.6)

This is a definition that’s worth reading more than once, and spending some time to digest it. Note that it takes several parameters:

- $T$ is the limit on the number of lines of the program $P$ that the generator needs to “fool”. The larger $T$ is, the stronger the generator.
- $\epsilon$ is how close is the output of the pseudorandom generator to the true uniform distribution over $\{0, 1\}^m$. The smaller $\epsilon$ is, the stronger the generator.
- $\ell$ is the input length and $m$ is the output length.

If $\ell \geq m$ then it is trivial to come up with such a generator: on input $s \in \{0, 1\}^\ell$, we can output $s_0, \ldots, s_{m-1}$. In this case $\Pr_{s \sim \{0, 1\}^\ell} [P(G(s)) = 1]$ will simply equal $\Pr_{r \sim \{0, 1\}^m} [P(r) = 1]$, no matter how many lines $P$ has. So, the smaller $\ell$ is and the larger $m$ is, the stronger the generator, and to get anything non-trivial, we need $m > \ell$.

Furthermore note that although our eventual goal is to fool probabilistic randomized algorithms that take an unbounded number of inputs, Definition 19.9 refers to finite and deterministic NAND-CIRC programs.

We can think of a pseudorandom generator as a “randomness amplifier.” It takes an input $s$ of $\ell$ bits chosen at random and expands these $\ell$ bits into an output $r$ of $m > \ell$ pseudorandom bits. If $\epsilon$ is small enough then the pseudorandom bits will “look random” to any NAND-CIRC program that is not too big. Still, there are two questions we haven’t answered:

- **What reason do we have to believe that pseudorandom generators with non-trivial parameters exist?**

- **Even if they do exist, why would such generators be useful to derandomize randomized algorithms?** After all, Definition 19.9 does not involve RNAND-TM or RNAND-RAM programs, but rather deterministic NAND-CIRC programs with no randomness and no loops.
We will now (partially) answer both questions. For the first question, let us come clean and confess we do not know how to prove that interesting pseudorandom generators exist. By interesting we mean pseudorandom generators that satisfy that \( \epsilon \) is some small constant (say \( \epsilon < 1/3 \)), \( m > \ell \), and the function \( G \) itself can be computed in \( \text{poly}(m) \) time. Nevertheless, Lemma 19.12 (whose statement and proof is deferred to the end of this chapter) shows that if we only drop the last condition (polynomial-time computability), then there do in fact exist pseudorandom generators where \( m \) is exponentially larger than \( \ell \).

At this point you might want to skip ahead and look at the statement of Lemma 19.12. However, since its proof is somewhat subtle, I recommend you defer reading it until you’ve finished reading the rest of this chapter.

19.3.2 From existence to constructivity
The fact that there exists a pseudorandom generator does not mean that there is one that can be efficiently computed. However, it turns out that we can turn complexity “on its head” and use the assumed non existence of fast algorithms for problems such as 3SAT to obtain pseudorandom generators that can then be used to transform randomized algorithms into deterministic ones. This is known as the Hardness vs Randomness paradigm. A number of results along those lines, most of which are outside the scope of this course, have led researchers to believe the following conjecture:

**Optimal PRG conjecture:** There is a polynomial-time computable function \( \text{PRG} : \{0, 1\}^* \rightarrow \{0, 1\} \) that yields an exponentially secure pseudorandom generator.

Specifically, there exists a constant \( \delta > 0 \) such that for every \( \ell \) and \( m < 2^{\delta \ell} \), if we define \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \) as \( G(s)_i = \text{PRG}(s, i) \) for every \( s \in \{0, 1\}^\ell \) and \( i \in [m] \), then \( G \) is a \((2^{\delta \ell}, 2^{-\delta \ell})\) pseudorandom generator.

The “optimal PRG conjecture” is worth while reading more than once. What it posits is that we can obtain \((T, \epsilon)\) pseudorandom generator \( G \) such that every output bit of \( G \) can be computed in time polynomial in the length \( \ell \) of the input, where \( T \) is exponentially large in \( \ell \) and \( \epsilon \) is exponentially small in \( \ell \). (Note that
A pseudorandom generator of the form we posit, where each output bit can be computed individually in time polynomial in the seed length, is commonly known as a pseudorandom function generator. For more on the many interesting results and connections in the study of pseudorandomness, see this monograph of Salil Vadhan.

We emphasize again that the optimal PRG conjecture is, as its name implies, a conjecture, and we still do not know how to prove it. In particular, it is stronger than the conjecture that $P \neq NP$. But we do have some evidence for its truth. There is a spectrum of different types of pseudorandom generators, and there are weaker assumptions than the optimal PRG conjecture that suffice to prove that $BPP = P$. In particular this is known to hold under the assumption that there exists a function $F \in \text{TIME}(2^{O(n)})$ and $\epsilon > 0$ such that for every sufficiently large $n$, $F_{\mid n}$ is not in $\text{SIZE}(2^{\epsilon n})$. The name “Optimal PRG conjecture” is non standard. This conjecture is sometimes known in the literature as the existence of exponentially strong pseudorandom functions.\textsuperscript{5}

19.3.3 Usefulness of pseudorandom generators

We now show that optimal pseudorandom generators are indeed very useful, by proving the following theorem:

**Theorem 19.10 — Derandomization of BPP.** Suppose that the optimal PRG conjecture is true. Then $BPP = P$.

**Proof Idea:**

The optimal PRG conjecture tells us that we can achieve exponential expansion of $\ell$ truly random coins into as many as $2^{\delta \ell}$ “pseudorandom coins.” Looked at from the other direction, it allows us to reduce the need for randomness by taking an algorithm that uses $m$ coins and converting it into an algorithm that only uses $O(\log m)$ coins. Now an algorithm of the latter type by can be made fully deterministic by enu-
merating over all the $2^{O(\log m)}$ (which is polynomial in $m$) possibilities for its random choices.

* We now proceed with the proof details.

**Proof of Theorem 19.10.** Let $F \in \text{BPP}$ and let $P$ be a NAND-TM program and $a, b, c, d$ constants such that for every $x \in \{0, 1\}^n$, $P(x)$ runs in at most $e \cdot n^d$ steps and $\Pr_{r \sim \{0,1\}^m}[P(x; r) = F(x)] \geq 2/3$. By “unrolling the loop” and hardwiring the input $x$, we can obtain for every input $x \in \{0, 1\}^n$ a NAND-CIRC program $Q_x$ of at most, say, $T = 10e \cdot n^d$ lines, that takes $m$ bits of input and such that $Q(r) = P(x; r)$.

Now suppose that $G : \{0, 1\}^\ell \rightarrow \{0, 1\}$ is a $(T, 0.1)$ pseudorandom generator. Then we could deterministically estimate the probability $p(x) = \Pr_{s \sim \{0,1\}^\ell}[Q_x(G(s)) = 1]$ up to 0.1 accuracy in time $O(T \cdot 2^\ell \cdot m \cdot \text{cost}(G))$ where $\text{cost}(G)$ is the time that it takes to compute a single output bit of $G$.

The reason is that we know that $\tilde{p}(x) = \Pr_{s \sim \{0,1\}^\ell}[Q_x(s) = 1]$ will give us such an estimate for $p(x)$, and we can compute the probability $\tilde{p}(x)$ by simply trying all $2^\ell$ possibilities for $s$. Now, under the optimal PRG conjecture we can set $T = 2^\delta T$ or equivalently $\ell = 1/\delta \log T$, and our total computation time is polynomial in $2^\ell = T^{1/\delta}$. Since $T \leq 10e \cdot n^d$, this running time will be polynomial in $n$.

This completes the proof, since we are guaranteed that $\Pr_{r \sim \{0,1\}^m}[Q_x(r) = F(x)] \geq 2/3$, and hence estimating the probability $p(x)$ to within 0.1 accuracy is sufficient to compute $F(x)$.

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### 19.4 $P = \text{NP AND BPP VS P}$

Two computational complexity questions that we cannot settle are:

- **Is $P = \text{NP}$?** Where we believe the answer is **negative**.
- **Is $\text{BPP} = P$?** Where we believe the answer is **positive**.

However we can say that the “conventional wisdom” is correct on at least one of these questions. Namely, if we’re wrong on the first count, then we’ll be right on the second one:

**Theorem 19.11 — Sipser–Gács Theorem.** If $P = \text{NP}$ then $\text{BPP} = P$.

Before reading the proof, it is instructive to think why this result is not “obvious.” If $P = \text{NP}$ then given any randomized algorithm $A$ and input $x$, ...
we will be able to figure out in polynomial time if there is a string \( r \in \{0, 1\}^m \) of random coins for \( A \) such that \( A(xr) = 1 \). The problem is that even if \( \Pr_{r \in \{0, 1\}^m}[A(xr) = F(x)] \geq 0.9999 \), it can still be the case that even when \( F(x) = 0 \) there exists a string \( r \) such that \( A(xr) = 1 \).

The proof is rather subtle. It is much more important that you understand the statement of the theorem than that you follow all the details of the proof.

**Proof Idea:**

The construction follows the “quantifier elimination” idea which we have seen in Theorem 15.5. We will show that for every \( F \in \text{BPP} \), we can reduce the question of some input \( x \) satisfies \( F(x) = 1 \) to the question of whether a formula of the form \( \exists u \in \{0, 1\}^m \forall v \in \{0, 1\}^n \enspace P(u, v) \) is true, where \( m, k \) are polynomial in the length of \( x \) and \( P \) is polynomial-time computable. By Theorem 15.5, if \( P = \text{NP} \) then we can decide in polynomial time whether such a formula is true or false.

The idea behind this construction is that using amplification we can obtain a randomized algorithm \( A \) for computing \( F \) using \( m \) coins such that for every \( x \in \{0, 1\}^n \), if \( F(x) = 0 \) then the set \( S \subseteq \{0, 1\}^m \) of coins that make \( A \) output 1 is extremely tiny, and if \( F(x) = 1 \) then it is very large. Now in the case \( F(x) = 1 \), one can show that this means that there exists a small number \( k \) of “shifts” \( s_0, \ldots, s_{k-1} \) such that the union of the sets \( S \oplus s_i \) (i.e., sets of the form \( \{ s \oplus s_i \mid s \in S \} \)) covers \( \{0, 1\}^m \), while in the case \( F(x) = 0 \) this union will always be of size at most \( k|S| \) which is much smaller than \( 2^m \). We can express the condition that there exists \( s_0, \ldots, s_{k-1} \) such that \( \bigcup_{i \in [k]} (S \oplus s_i) = \{0, 1\}^m \) as a statement with a constant number of quantifiers.

* \( 1 \)

**Proof of Theorem 19.11.** Let \( F \in \text{BPP} \). Using Theorem 19.4, there exists a polynomial-time algorithm \( A \) such that for every \( x \in \{0, 1\}^n \),

\[
\Pr_{x \in \{0, 1\}^n}[A(xr) = F(x)] \geq 1 - 2^{-n}
\]

where \( m \) is polynomial in \( n \). In particular (since an exponential dominates a polynomial, and we can always assume \( n \) is sufficiently large), it holds that

\[
\Pr_{x \in \{0, 1\}^n}[A(xr) = F(x)] \geq 1 - \frac{1}{10^m}. \tag{19.7}
\]

Let \( x \in \{0, 1\}^n \), and let \( S_x \subseteq \{0, 1\}^m \) be the set \( \{ r \in \{0, 1\}^m : A(xr) = 1 \} \). By our assumption, if \( F(x) = 0 \) then \( |S_x| \leq \frac{1}{10^m} 2^m \) and if \( F(x) = 1 \) then \( |S_x| \geq (1 - \frac{1}{10^m}) 2^m \).

For a set \( S \subseteq \{0, 1\}^m \) and a string \( s \in \{0, 1\}^m \), we define the set \( S \oplus s \) to be \( \{ r \oplus s : r \in S \} \), where \( \oplus \) denotes the XOR operation. That is, \( S \oplus s \) is the set \( S \) “shifted” by \( s \). Note that \( |S \oplus s| = |S| \). (Please make sure that you see why this is true.)

Figure 19.7: If \( F \in \text{BPP} \) then through amplification we can ensure that there is an algorithm \( A \) to compute \( F \) on \( n \)-length inputs and using \( m \) coins such that \( \Pr_{r \in \{0, 1\}^m}[A(xr) \neq F(x)] \ll 1/poly(m) \). Hence if \( F(x) = 1 \) then almost all of the \( 2^m \) choices for \( r \) will cause \( A(xr) \) to output 1, while if \( F(x) = 0 \) then \( A(xr) = 0 \) for almost all \( r \)’s. To prove the Sipser–Gács Theorem we consider several “shifts” of the set \( S \subseteq \{0, 1\}^m \) of the coins \( r \) such that \( A(xr) = 1 \). If \( F(x) = 1 \) then we can find a set of \( k \) shifts \( s_0, \ldots, s_{k-1} \) for which \( \bigcup_{i \in [k]} (S \oplus s_i) = \{0, 1\}^m \). If \( F(x) = 0 \) then for every such set \( \bigcup_{i \in [k]} |S_{i} \mid \leq k|S| \ll 2^m \). We can phrase the question of whether there is such a set of shift using a constant number of quantifiers, and so can solve it in polynomial time if \( P = \text{NP} \).
The heart of the proof is the following two claims:

CLAIM I: For every subset $S \subseteq \{0, 1\}^m$, if $|S| \leq \frac{1}{1000m^2}$, then for every $s_0, \ldots, s_{100m-1} \in \{0, 1\}^m$, $\bigcup_{i \in [100m]} (S \oplus s_i) \subseteq \{0, 1\}^m$.

CLAIM II: For every subset $S \subseteq \{0, 1\}^m$, if $|S| \geq \frac{1}{2^m}$ then there exist a set of string $s_0, \ldots, s_{100m-1}$ such that $\bigcup_{i \in [100m]} (S \oplus s_i) = \{0, 1\}^m$.

CLAIM I and CLAIM II together imply the theorem. Indeed, they mean that under our assumptions, for every $x \in \{0, 1\}^n$, $F(x) = 1$ if and only if

$$\exists s_0, \ldots, s_{100m-1} \in \{0, 1\}^m \cup \bigcup_{i \in [100m]} (S \oplus s_i) = \{0, 1\}^m$$  \hspace{1cm} (19.8)

which we can re-write as

$$\exists s_0, \ldots, s_{100m-1} \in \{0, 1\}^m \forall w \in \{0, 1\}^m (w \in (S \oplus s_0) \lor w \in (S \oplus s_1) \lor \ldots \lor w \in (S \oplus s_{100m-1}))$$  \hspace{1cm} (19.9)

or equivalently

$$\exists s_0, \ldots, s_{100m-1} \in \{0, 1\}^m \forall w \in \{0, 1\}^m \left( A(x(w \oplus s_0)) = 1 \lor A(x(w \oplus s_1)) = 1 \lor \ldots \lor A(x(w \oplus s_{100m-1})) = 1 \right)$$  \hspace{1cm} (19.10)

which (since $A$ is computable in polynomial time) is exactly the type of statement shown in Theorem 15.5 to be decidable in polynomial time if $P = NP$.

We see that all that is left is to prove CLAIM I and CLAIM II.

CLAIM I follows immediately from the fact that

$$\bigcup_{i \in [100m-1]} |S \oplus s_i| \leq \sum_{i=0}^{100m-1} |S \oplus s_i| = \sum_{i=0}^{100m-1} |S_x| = 100m |S_x|.$$  \hspace{1cm} (19.11)

To prove CLAIM II, we will use a technique known as the probabilistic method (see the proof of Lemma 19.12 for a more extensive discussion). Note that this is a completely different use of probability than in the theorem statement, we just use the methods of probability to prove an existential statement.

Proof of CLAIM II: Let $S \subseteq \{0, 1\}^m$ with $|S| \geq 0.5 \cdot 2^m$ be as in the claim’s statement. Consider the following probabilistic experiment: we choose $100m$ random shifts $s_0, \ldots, s_{100m-1}$ independently at random in $\{0, 1\}^m$, and consider the event $GOOD$ that $\bigcup_{i \in [100m]} (S \oplus s_i) = \{0, 1\}^m$. To prove CLAIM II it is enough to show that $\Pr[GOOD] > 0$, since that means that in particular there must exist shifts $s_0, \ldots, s_{100m-1}$ that satisfy this condition.

For every $z \in \{0, 1\}^m$, define the event $BAD_z$ to hold if $z \notin \bigcup_{i \in [100m]} (S \oplus s_i)$. The event $GOOD$ holds if $BAD_z$ fails for every $z \in \{0, 1\}^m$, and so our goal is to prove that $\Pr[\bigcup_{z \in \{0, 1\}^m} BAD_z] < 1$. By
the union bound, to show this, it is enough to show that \( \Pr[BAD_z] < 2^{-m} \) for every \( z \in \{0, 1\}^m \). Define the event \( BAD_z \) to hold if \( z \notin S \oplus s_i \).

Since every shift \( s_i \) is chosen independently, for every fixed \( z \) the events \( BAD_z, \ldots, BAD_z^{100m-1} \) are mutually independent,\(^6\) and hence

\[
\Pr[BAD_z] = \Pr[\bigcap_{i \in [100m-1]} BAD_z^i] = \prod_{i=0}^{100m-1} \Pr[BAD_z^i]. \quad (19.12)
\]

So this means that the result will follow by showing that \( \Pr[BAD_z^i] \leq \frac{1}{2} \) for every \( z \in \{0, 1\}^m \) and \( i \in [100m] \) (as that would allow to bound the righthand side of (19.12) by \( 2^{-100m} \)). In other words, we need to show that for every \( z \in \{0, 1\}^m \) and set \( S \subseteq \{0, 1\}^m \) with \( |S| \geq \frac{1}{2} 2^m \),

\[
\Pr_{s \in \{0,1\}^m} \left[ z \in S \oplus s \right] \geq \frac{1}{2}. \quad (19.13)
\]

To show this, we observe that \( z \in S \oplus s \) if and only if \( s \in S \oplus z \) (can you see why). Hence we can rewrite the probability on the lefthand side of (19.13) as \( \Pr_{s \in \{0,1\}^m} \left[ s \in S \oplus z \right] \) which simply equals \( |S \oplus z|/2^m = |S|/2^m \geq 1/2! \) This concludes the proof of CLAIM I and hence of Theorem 19.11.

\[\blacksquare\]

### 19.5 NON-CONSTRUCTIVE EXISTENCE OF PSEUDORANDOM GENERATORS (ADVANCED, OPTIONAL)

We now show that, if we don’t insist on constructivity of pseudorandom generators, then we can show that there exists pseudorandom generators with output that exponentially larger in the input length.

**Lemma 19.12 — Existence of inefficient pseudorandom generators.** There is some absolute constant \( C \) such that for every \( \epsilon, T, \) if \( \ell > C(\log T + \log(1/\epsilon)) \) and \( m \leq T \), then there is an \((T, \epsilon)\) pseudorandom generator \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \).

**Proof Idea:**

The proof uses an extremely useful technique known as the “probabilistic method” which is not too hard mathematically but can be confusing at first.\(^7\) The idea is to give a “non constructive” proof of existence of the pseudorandom generator \( G \) by showing that if \( G \) was chosen at random, then the probability that it would be a valid \((T, \epsilon)\) pseudorandom generator is positive. In particular this means that there exists a single \( G \) that is a valid \((T, \epsilon)\) pseudorandom generator. The probabilistic method is just a proof technique to demonstrate the existence of such a function. Ultimately, our goal is to show the existence of a deterministic function \( G \) that satisfies the condition.

\[\text{There is a whole (highly recommended) book by Alon and Spencer devoted to this method.}\]
The above discussion might be rather abstract at this point, but would become clearer after seeing the proof.

**Proof of Lemma 19.12.** Let \( \epsilon, T, \ell, m \) be as in the lemma’s statement. We need to show that there exists a function \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \) that “fools” every \( T \) line program \( P \) in the sense of (19.6). We will show that this follows from the following claim:

**Claim I:** For every fixed NAND-CIRC program \( P \), if we pick \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \) at random then the probability that (19.6) is violated is at most \( 2^{-T^2} \).

Before proving Claim I, let us see why it implies Lemma 19.12. We can identify a function \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \) with its “truth table” or simply the list of evaluations on all its possible \( 2^\ell \) inputs. Since each output is an \( m \) bit string, we can also think of \( G \) as a string in \( \{0, 1\}^{m \cdot 2^\ell} \). We define \( \mathcal{F}_m^{\ell} \) to be the set of all functions from \( \{0, 1\}^\ell \) to \( \{0, 1\}^m \). As discussed above we can identify \( \mathcal{F}_m^{\ell} \) with \( \{0, 1\}^{m \cdot 2^\ell} \) and choosing a random function \( G \sim \mathcal{F}_m^{\ell} \) corresponds to choosing a random \( m \cdot 2^{\ell} \)-long bit string.

For every NAND-CIRC program \( P \) let \( B_P \) be the event that, if we choose \( G \) at random from \( \mathcal{F}_m^{\ell} \) then (19.6) is violated with respect to the program \( P \). It is important to understand what is the sample space that the event \( B_P \) is defined over, namely this event depends on the choice of \( G \) and so \( B_P \) is a subset of \( \mathcal{F}_m^{\ell} \). An equivalent way to define the event \( B_P \) is that it is the subset of all functions mapping \( \{0, 1\}^\ell \) to \( \{0, 1\}^m \) that violate (19.6), or in other words:

\[
B_P = \left\{ G \in \mathcal{F}_m^{\ell} \mid \frac{1}{2^\ell} \sum_{s \in \{0, 1\}^\ell} P(G(s)) - \frac{1}{2^m} \sum_{r \in \{0, 1\}^m} P(r) > \epsilon \right\}
\]

(We’ve replaced here the probability statements in (19.6) with the equivalent sums so as to reduce confusion as to what is the sample space that \( B_P \) is defined over.)

To understand this proof it is crucial that you pause here and see how the definition of \( B_P \) above corresponds to (19.14). This may well take re-reading the above text once or twice, but it is a good exercise at parsing probabilistic statements and learning how to identify the sample space that these statements correspond to.

Now, we’ve shown in Theorem 5.9 that up to renaming variables (which makes no difference to program’s functionality) there are \( 2^{O(T \log T)} \) NAND-CIRC programs of at most \( T \) lines. Since \( T \log T < T^2 \) for sufficiently large \( T \), this means that if Claim I is true, then by the union bound it holds that the probability of the union of
\( B_P \) over all NAND-CIRC programs of at most \( T \) lines is at most
\[ 2^{O(T \log T)}2^{-T^2} < 0.1 \] for sufficiently large \( T \). What is important for us about the number 0.1 is that it is smaller than 1. In particular this means that there exists a single \( G^* \in \mathcal{F}^m \) such that \( G^* \) does not violate (19.6) with respect to any NAND-CIRC program of at most \( T \) lines, but that precisely means that \( G^* \) is a \((T, \epsilon)\) pseudorandom generator.

Hence to conclude the proof of Lemma 19.12, it suffices to prove Claim I. Choosing a random \( G : \{0, 1\}^\ell \rightarrow \{0, 1\}^m \) amounts to choosing \( L = 2^\ell \) random strings \( y_0, \ldots, y_{L-1} \in \{0, 1\}^m \) and letting \( G(x) = y_x \) (identifying \( \{0, 1\}^\ell \) and \( [L] \) via the binary representation). This means that proving the claim amounts to showing that for every fixed function \( P : \{0, 1\}^m \rightarrow \{0, 1\} \), if \( L > 2^{C\left(\log T + \log \epsilon\right)} \) (which by setting \( C > 4 \), we can ensure is larger than \( 10T^2/\epsilon^2 \)) then the probability that

\[ \left| \frac{1}{L} \sum_{i=0}^{L-1} P(y_i) - \Pr_{y \sim \{0,1\}^m}[P(y) = 1] \right| > \epsilon \] (19.15)

is at most \( 2^{-T^2} \).

(??) follows directly from the Chernoff bound. Indeed, if we let for every \( i \in [L] \) the random variable \( X_i \) denote \( P(y_i) \), then since \( y_0, \ldots, y_{L-1} \) is chosen independently at random, these are independently and identically distributed random variables with mean

\[ \mathbb{E}_{y \sim \{0,1\}^m}[P(y)] = \Pr_{y \sim \{0,1\}^m}[P(y) = 1] \]

and hence the probability that they deviate from their expectation by \( \epsilon \) is at most \( 2 \cdot 2^{-L/2} \).

\( \blacksquare \)

Lecture Recap

- We can model randomized algorithms by either adding a special “coin toss” operation or assuming an extra randomly chosen input.
- The class \( \text{BPP} \) contains the set of Boolean functions that can be computed by polynomial time randomized algorithms.
- We know that \( P \subseteq \text{BPP} \subseteq \text{EXP} \).
- We also know that \( \text{BPP} \subseteq P_{/\text{poly}} \).
- The relation between \( \text{BPP} \) and \( \text{NP} \) is not known, but we do know that if \( P = \text{NP} \) then \( \text{BPP} = P \).
- Pseudorandom generators are objects that take a short random “seed” and expand it to a much longer output that “appears random” for efficient algorithms. We conjecture that exponentially strong pseudorandom generators exist. Under this conjecture, \( \text{BPP} = P \).
19.6 EXERCISES

R Remark 19.13 — Disclaimer. Most of the exercises have been written in the summer of 2018 and haven’t yet been fully debugged. While I would prefer people do not post online solutions to the exercises, I would greatly appreciate if you let me know of any bugs. You can do so by posting a GitHub issue about the exercise, and optionally complement this with an email to me with more details about the attempted solution.

19.7 BIBLIOGRAPHICAL NOTES

19.8 FURTHER EXPLORATIONS

Some topics related to this chapter that might be accessible to advanced students include: (to be completed)

19.9 ACKNOWLEDGEMENTS
V

ADVANCED TOPICS